## Bayesian Approaches I

## Bayesian

- data are treated as fixed observations
- models (parameters) are treated as random variables
- we compute the probability of all models
- we end up with a richer understanding of relative probability of all models


## Frequentist

- data (sample) treated as a random variable
- models (population parameters) are treated as fixed quantities
- we compute the probability of one model (H0)
- we make a decision (reject H0 or not)


## Bayes Theorem

$$
p(B \mid A)=\frac{p(A \mid B) p(B)}{p(A)}
$$

- probability of B, given A, equals probability of A, given B, times prob of B, divided by probability of $A$
- $p(B \mid A)$ is the posterior
- $p(A \mid B)$ is the likelihood
- $p(B)$ is the prior
- $p(A)$ is the evidence


# Bayes Theorem 

$$
p(\text { model } \mid d a t a)=\frac{p\left(\text { data } a \left\lvert\, \begin{array}{c}
\text { likelihood } \\
\text { model }
\end{array}\right.\right) p\binom{\text { prior }}{\text { model }}}{p(\text { data }) \text { evidence }}
$$

- probability of model, given data, equals probability of data, given model, times prob of model, divided by probability of data
- P (model|data) is the posterior
- $p$ (data|model) is the likelihood
- $p$ (model) is the prior
- $p$ (data) is the evidence


## Bayes Theorem

$$
p(\text { model } \mid \text { data })=\frac{p(\text { data } \mid \text { model }) p(\text { model })}{p(\text { data })}
$$

- p (data), the marginal probability of data across all models, can be computed as the sum of conditional probabilities of data given each model:

$$
p(d a t a)=\sum_{\text {model }_{i}} p\left(\text { data }^{2} \text { model }_{i}\right) p\left(\text { model }_{i}\right)
$$

## Bayes Theorem

$$
p(\text { model } \mid \text { data })=\frac{p(\text { data } \mid \text { model }) p(\text { model })}{p(\text { data })}
$$

- We will first look at a discrete probability example, using single-point probabilities, to show how these calculations work
- We will then look at an example of this approach using continuous probability distributions instead of point probabilities


## Discrete Example

- Let's say you take a home pregnancy test and it comes out positive. What is the probability that you are pregnant?
- Let's say we know the test is $90 \%$ accurate
- The "data" we have is
- $p$ (test+ | preg) $=0.90$
- and test was +
- We want the prob of the "model": preg
- we want to know p(preg | test+)


## Discrete Example

$$
\begin{aligned}
p(B \mid A) & =\frac{p(A \mid B) p(B)}{p(A)} \\
p(\text { preg } \mid \text { test }+) & =\frac{p(\text { test }+\mid \text { preg }) p(\text { preg })}{p(\text { test }+)}
\end{aligned}
$$

- $\mathrm{p}($ test $+\mid$ preg $)=0.90$ (accuracy of test)
- $p$ (test+ $\mid$ not $p r e g)=0.50$ (false pos rate)
- let's say we also estimate $p(p r e g) \sim 0.15$


## Discrete Example $p($ preg $\mid$ test +$)=\frac{p(\text { test }+\mid \text { preg }) p(\text { preg })}{p(\text { test }+) ?}$

- $p($ test $+\mid$ preg $)=0.90$
- $p($ preg $)=0.15$
- what is $\mathrm{P}($ test +$) ? \quad p\left(\right.$ data $\left.^{2}\right)=\sum_{\text {model }_{i}} p\left(\right.$ data $^{\prime} \mid$ model $\left._{i}\right) p\left(\right.$ model $\left._{i}\right)$
- $p($ test +$)=p($ test $+\mid$ preg $) p($ preg $)+p($ test $+\mid$ notpreg $) p($ notpreg $)$
- $\mathrm{p}($ test + ) $=(.90)(.15)+(.50)(.85)=0.56$

> Discrete Example $p($ preg|test +$)=\frac{p(\text { test }+\mid \text { preg })(\text { (preg })}{p(\text { test }+) .56}$

- $p($ test +$)=(.90)(.15)+(.50)(.85)=.56$
- so $\mathrm{p}($ preg $\mid$ test +$)=(.90)(.15) /(.56)=.241$
- so probability of pregnant given pos test is 24.1\%

$$
\begin{gathered}
\text { Effect of the prior } \\
p(\text { model } \mid \text { data })=\frac{p(\text { datatalibod } \text { model }) \text { p(prior } \text { model })}{p(\text { data })}
\end{gathered}
$$

- the posterior $p$ (model | data) is proportional to the likelihood $p$ (data | model) multiplied by the prior $p$ (model)
- our prior expectation (or previous findings, i.e. data) modulates our prediction of the future
- can be viewed as both a virtue and a shortcoming of the Bayesian approach


## Effect of the prior

$$
p(\text { preg } \mid \text { test }+)=\frac{\stackrel{.90}{p^{2}(\text { test }+\mid \text { preg }) p(\text { preg })}}{p(\text { test }+) .56}
$$

| prior | posterior |
| :---: | :---: |
| 0.1 | 0.17 |
| 0.2 | 0.31 |
| 0.3 | 0.44 |
| 0.4 | 0.55 |
| 0.5 | 0.64 |


| prior | posterior |
| :---: | :---: |
| 0.6 | 0.73 |
| 0.7 | 0.81 |
| 0.8 | 0.88 |
| 0.9 | 0.94 |
| 0.99 | 0.99 |

## Updating the Model $p($ model $\mid$ data $)=\frac{p(\text { data } a \mid \text { ilelihood } \text { model }) p\left(\begin{array}{c}\text { prior } \\ \text { model })\end{array}\right.}{p(\text { data })}$

- When you collect new data, you can update your model
- the posterior from the previous model becomes the prior for the new model


## Updating the Model

$$
p(\text { model } \mid \text { data })=\frac{p(\text { datalalihood } \text { model }) p(\text { morior } \text { model })}{p(\text { data })}
$$

- Let's say you take another preg test
- We know from our previous calculation:
- $\mathrm{P}($ preg $\mid$ test + ) $=.241$
- The other quantities are the same
- p(test+ | preg) $=0.90$ (accuracy of test)
- $p($ test $+\mid$ not preg $)=0.50$ (false pos rate)


## Updating the Model

- Let's say you take another preg test
- We know from our previous calculation:
- $\mathrm{P}($ preg $\mid$ test + ) $=.241$
- this becomes our new prior
$p(p r e g \mid t e s t+)=\frac{p(t e s t+\mid p r e g) p(p r e g)}{p(t e s t+)}$
$p(p r e g \mid t e s t+)=\frac{p(t e s t+\mid p r e g) p(p r e g)}{p(t e s t+\mid p r e g) p(p r e g)+p(t e s t+\mid n o t p r e g) p(n o t p r e g)}$
- $p($ preg $\mid$ test + ) $=$
$(.90)(.24 I) /((.90)(.24 I)+(.50)(I-.24 I))$
$=.364$
- so after a second positive test, p (preg | test+) is now 36.4\%

| Test \# | Test 3 | Test 4 | Test 5 | Test 6 | Test 7 | Test 8 | Test 9 | Test 10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| p(pregltest+) | 0.51 | 0.65 | 0.77 | 0.86 | 0.92 | 0.95 | 0.97 | 0.98 |

## Updating the Model

- seems like an appropriate thing to do in science
- when new data are gathered, we can reevaluate a hypothesis
- we do not begin anew (ignorant) each time we ask a question
- previous research provides us information about the merits of the hypothesis


## Bayes with Distributions

- in previous example, the likelihood and prior were both single quantities (point probabilities)
- typically Bayesian approaches use full probability distributions
- essentially allows us to evaluate probability of a whole range of possible models, at once


## Bayes with Distributions

- don't worry, remember probability distributions are just mathematical functions of a parameter vector

$$
p(k \mid n, p)=\frac{n!}{k!(n-k)!} p^{k}(1-p)^{n-k}
$$

- e.g. binomal prob of $k$ successes in $n$ trials with prob(success) $p$, is

 mean mu and standard dev sigma, is


## Bayes with Distributions

$$
p(\text { model } \mid \text { data })=\frac{p(\text { datalalihood } \text { model }) p(\text { morior } \text { model })}{p(\text { data })_{\text {evidence }}^{\text {pil }}}
$$

- let's look at an example: coin flipping
- is my coin fair?
- "data" are 3 flips of the coin: (H, H,T)
- "model" is a proposed process by which the outcome of our coin flip is determined


## Bayes with Distributions

$$
p(\text { model } \mid \text { data })=\frac{p(\text { data } \mid \stackrel{\text { likelihood }}{\text { model }}) p(\stackrel{\text { prior }}{\text { model }})}{p(\text { data }) \text { evidence }}
$$

- since outcomes are binary ( $\mathrm{H}, \mathrm{T}$ ) a natural choice of model is a binomial distribution
- we know the likelihood function for a binomial model is

$$
p(k \mid n, p)=\frac{n!}{k!(n-k)!} p^{k}(1-p)^{n-k}
$$

- so our "data" are: $\mathrm{n}=3$ tries, $\mathrm{k}=2$ successes (assume Heads=success,Tails=failure)
- likelihood function gives us $p(k \mid n, p)$ but what we want is the posterior: $p(p \mid n, k)$ where $p$ is prob(success) (fair is $p=0.50$ )
- according to Bayes theorem this equals likelihood*prior/evidence


## Bayes with Distributions

$$
p(\text { model } \mid d a t a)=\frac{p\left(\text { data } \left\lvert\, \begin{array}{c}
\text { likelihood }
\end{array} \stackrel{\stackrel{\text { prior }}{\text { model }}) p(\text { model })}{\text { mod }}\right.\right)}{p(\text { data }) \text { evidence }}
$$

- what should our prior be?
- prior is probability of "model" == probability distribution over possible values of $p$
- we could decide on an "uninformative prior", postulating that all values of $p$ are equally likely:



## Bayes with Distributions

$$
p(\text { model } \mid d a t a)=\frac{p(\text { data } \mid \stackrel{\text { likelihood }}{\text { model }}) p(\stackrel{\text { prior }}{\text { model }})}{p(\text { data }) \text { evidence }}
$$

- we could decide that since every coin we have seen in the past has been fair, we expect that this coin will be fair as well and so $p$ will likely be $=0.50$
- but how unlikely are values other than p ?
- very unlikely?
- moderately unlikely?
- not terribly unlikely but stillmess likely than .50 ?



## Bayes with Distributions

$$
p(\text { model } \mid \text { data })=\frac{p\left(\text { data } \left\lvert\, \begin{array}{c}
\text { likelihood } \\
\text { model }
\end{array}\right.\right) p\binom{\text { prior }}{\text { model }}}{p(\text { data }) \text { evidence }}
$$

- it's totally up to us to decide on the prior, in several aspects:
I. scientific/theoretical Q : in general, what should its shape be?

2. practical Q : how do I characterize the prior?

- "by hand", e.g. as a table (a list) of parameter values \& probabilities
- "algebraically", as a mathematical equation
A. any old function of our choosing, OR
B. a specific equation that will help us later in computing the posterior (known as a conjugate prior)


## Bayes with Distributions

- Two general approaches to computing the posterior:
- Analytic: choosing a likelihood model and a conjugate prior from a (relatively short) list of known forms, and taking advantage of clever algebra/calculus that results in a very simple expression for the posterior


## Bayes with Distributions

- Numerical: you're free to specify your likelihood and your prior as whatever you want, and use iterative computing methods and powerful computers to estimate the posterior distribution
- grid approximation approach
- Markov Chain Monte-Carlo (MCMC)


## Analytic Approach

- recall our data: 3 coin flips, 2 successes (2 HEADS, one TAILS)
- is the coin fair? === what is prob $p$ in our binomial model
- we want the posterior: ${ }_{p(M \mid D)}=\frac{p(D \mid M) p(M)}{p(D)}$
- likelihood $P(D \mid M)$ is given by the binomial distribution

$$
p(k \mid n, p)=\frac{n!}{k!(n-k)!} p^{k}(1-p)^{n-k}
$$

- it turns out that a conjugate prior for the binomial, is the Beta distribution
- http://en.wikipedia.org/wiki/Conjugate_prior


## Conjugate Priors

- If the posterior distribution $p(\theta \mid x)$ is in the same family as the prior probability distribution $p(\theta)$, the prior and posterior are then called conjugate distributions, and the prior is called a conjugate prior for the likelihood


## The Beta Distribution

$$
f(x, \alpha, \beta)=\frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)}
$$

http://en.wikipedia.org/wiki/Beta_distribution

- crystal clear, right? :) no of course not
- don't fret though ... this is just a mathematical equation
- it takes in parameters alpha and beta and spits out nice looking curves for $x$ values between 0 and I
- this is convenient for characterizing prior on $p$, since in our coin, p is betwen 0 and I


## The Beta Distribution



Beta(5,20)


Beta(20,20)


Beta(20,5)



## Conjugate Priors

- when you use a prior that is a conjugate for the likelihood, then computing the posterior turns out to be a piece of cake
- clever calculus ninjas have worked out the algebra, and often the posterior can be expressed as a really simple manipulation of the parameters of the likelihood and prior


## Conjugate Priors

- for example for the binomial, we have our likelihood function $\operatorname{prob}(k \mid n, p)=\operatorname{binomial}(k, n, p)$
- and if we specify our prior using a Beta distribution $\operatorname{prob}(p)=\operatorname{beta}(\alpha, \beta)$
- then the posterior turns out to be equal to another Beta function, with modified alpha and beta parameters: $\operatorname{prob}(p \mid k, n)=\operatorname{beta}(k+\alpha, N-k+\beta)$
- thank you calculus ninjas!
- we don't even need to calculate anything


## Back to our example

- coin flip: $\mathrm{n}=3$ trials, $\mathrm{k}=2$ success
- likelihood is binomial(n,k,p)
- $\mathrm{n}=3, \mathrm{k}=2, \mathrm{p}$ is unknown
- prior is Beta(alpha,beta)
- let's choose a flat prior, alpha=I, beta=I
- our calculus ninjas gave us:
- posterior is $\operatorname{Beta}(2+\mathrm{I}, 3-2+\mathrm{I})$


## Our curvy posterior



- MLE of $p$ is 0.667
- the posterior also gives us the entire curve


## Describing the posterior

Beta(3,2)

- Graphically

- Summary statistics
- Analytic
- well known expressions for mean, variance, mode, MLE, etc...
- e.g. mean of a Beta dist is $\frac{\alpha}{\alpha+\beta}$
- variance is $\frac{\alpha \beta}{(\alpha+\beta)^{2}(\alpha+\beta+1)}$


## Describing the posterior

- Numerical
- use a random number generator to draw a large number of values from the posterior distribution, then compute summary stats from those random draws
- in programs like R we have a whole set of random number generators for lots of probability distributions
- normal, beta, binomial, exponential, poission, etc etc etc...


## Numerical Example

- let's compute the $95 \%$ credible interval sometimes called Highest Density Region (HDR)

```
> mysamp <- rbeta(10000, 3, 2)
> hist(mysamp, breaks=50, col="blue")
> ci95 = quantile(mysamp, c(.025,.975))
> ci95
    2.5% 97.5%
0.1940172 0.9335989
> abline(v=2/3, col="red", lwd=2, lty=2)
> abline(v=ci95, col="red", lwd=2)
```

- or: 50\% credible interval
> ci50 = quantile(mysamp, c(.25, .75))
> abline(v=ci50, col="orange", lwd=5)

Histogram of mysamp


## Numerical Example

- or any other quantity you could ever want
- after all, you have the ability to sample from the posterior distribution as much as you want

Histogram of mysamp

- i.e. you can sample from prob(model | data)
and characterize its entire shape, over the full range of possible values of the model
- essentially you can evaluate the relative prob of all models


## Try a different Prior

- We used a flat prior for the previous example

- Let's repeat but use a prior that expresses our evidence to date that coins are in fact fair Beta $(20,20)$

Beta $(20,20)$


## Try a different Prior

- likelihood
- flat prior didn't really change likelihood
- so posterior essentially equals likelihood



## Try a different Prior

- coin flip: $\mathrm{n}=3$ trials, $\mathrm{k}=2$ success
- likelihood is binomial(n,k,p)
- $\mathrm{n}=3, \mathrm{k}=2, \mathrm{p}$ is unknown
- prior is Beta(alpha,beta)
- let's choose an informative prior, alpha=20, beta=20
- our calculus ninjas gave us:
- posterior is $\operatorname{Beta}(2+20,3-2+20)$



## Effect of Prior



# Criticisms of Bayesian Approach 

- the prior: too much "subjectivity"?
- data fixed, models (parameters) random


## Advantages

- interval estimates (and other such measures of posterior) have a clearer meaning than Cls in frequentist approaches
- frequentist orientation around "repeated sampling" is unrealistic, we in fact only sample (do our experiment) once
- frequentist involves testing only one hypothesis (model) : the null hypothesis ... Bayesian estimates probability of all models (parameter values)
- in Bayesian approach we get full posterior distribution, a much richer picture than just a mean $+/-\mathrm{Cl}$ or s.e.
- Bayesian approach allows for incorporating previous findings in a principled way


## Next Class

- grid approximation approach (discretizing the prior)
- multidimensional models
- Markov Chain Monte Carlo (MCMC)


## gentle books

| Scott M. Lynch |
| :--- |
| Introduction |
| to Applied Bayesian |
| Statistics and Estimation for Social and |
| foriences |
| for Social Scientists |
| Springer |

Scott M. Lynch

## Introduction <br> to Applied Bayesian <br> Statistics and Estimation for Social Scientists

Q Springer

John K. Kruschke


A Tutorial with R and BUGS

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